

# Applied Statistics (BUGS)

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# BUGS

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## 4.2 Normal distribution

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## Section 4.2

```
# First we specify the statistical model
model
{
#  $time_i \sim N(\mu, \sigma^2)$ ,  $i = 1, \dots, n$ 
for(i in 1:n)
{
time[i] ~ dnorm(mu, inv.sigma2)
}
# Prior:  $\mu \sim N(0, 10^6)$ 
mu ~ dnorm(0, 0.000001)
# Prior:  $\sigma^{-2} \sim G(10^{-4}, 10^{-4})$ 
inv.sigma2 ~ dgamma(0.0001, 0.0001)
sigma2 <- 1/inv.sigma2
```

## Section 4.2

```
sigma <- sqrt(sigma2)
# the end of the model statement
}
# set initial values for MCMC
list(mu = 0, inv.sigma2 = 1)
# define variables for the dataset
list(n = 20)
# define variable time
time[]
182
...
365
END
```

## 2.4 Binomial distribution

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## Section 2.4

```
# First we specify the statistical model
model
{
#  $y \sim \text{Binomial}(n, p)$ 
y ~ dbin(p, n)
# Prior:  $p \sim \text{Beta}(a, b)$ 
p ~ dbeta(a, b)
#  $\text{pr.ge.half} = I(p > 0.5)$  to estimate  $Pr(p > 0.5|y)$ 
pr.ge.half <- step(p-0.5)
# Generate  $y.\text{pred} \sim \text{Binomial}(20, p)$  where  $p \sim \pi(p|y)$ 
y.pred ~ dbin(p, 20)
# the end of the model statement
}
```

## Section 2.4

```
# define variables for the dataset
list(n = 27, y = 11, a = 3.26, b = 7.18)
# set initial values for MCMC
list( p= 0.4, y.pred=5 )
```



## **3.3 Poisson distribution**

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## Section 3.3

```
# First we specify the statistical model
model
{
#  $y \sim \text{Poisson}(\lambda_y)$ ,  $\lambda_y = eo \times \lambda$ 
y ~ dpois(lambda.y)
lambda.y <- eo*lambda
# Prior:  $\lambda \sim G(\alpha, \beta)$ 
lambda ~ dgamma(alpha, beta)
# the end of the model statement
}
# define variables for the dataset
list( y=1, eo=66, alpha=16, beta=15174)
# set initial values for MCMC
list( lambda = 0.001)
```

## 6.9 Cauchy Distribution

---

## Section 6.9

```
# First we specify the statistical model
model
{
for(i in 1:n)
{
# Likelihood:  $lik_i = \left\{ 1 + \frac{(diff_i - \mu)^2}{\sigma^2} \right\}^{-1}$ 
lik[i] <-
1/(sigma*(1+((diff[i]-mu)/sigma)*((diff[i]-mu)/sigma)))
# 0's trick. Define  $\lambda_i = -\log(lik_i) + 10000 > 0$ 
lam[i] <- -log(lik[i])+10000
# Define an auxiliary variable  $zeros_i \equiv 0$ 
zeros[i] <- 0
```

## Section 6.9

```
#  $zeros_i \sim \text{Poisson}(\lambda_i)$ ,  $Pr(zeros_i = 0) = lik_i * \exp(-10000)$ 
zeros[i] ~ dpois(lam[i])
# Prior:  $\mu \sim N(0, 10^6)$ 
mu ~ dnorm(0, 0.000001)
# Prior:  $\sigma \sim G(10^{-3}, 10^{-3})$ 
sigma ~ dgamma(0.001, 0.001)
lsigma <- log(sigma)
# the end of the model statement
}
# set initial values for MCMC
list( mu = 0, sigma=1 )
# define variables for the dataset
list(n = 15)
```

## Section 6.9

```
# define variable diff
diff[]
-67
...
75
END
```

## **6.10 Heart Transplant Data**

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## Section 6.10

```
# First we specify the statistical model
model
{
  for(i in 1:n)
  {
    # Likelihood when  $state_i = 0$ ,  $transplant_i = 0$ 
    #  $\log(lik_i) = \log p + p \log \lambda - (p + 1) \log(\lambda + survtime_i)$ 
    llk_nt_0[i] <-
    log(p)+p*log(lambda)-(p+1)*log(lambda+survtime[i])
    # Likelihood when  $state_i = 1$ ,  $transplant_i = 0$ 
    #  $\log(lik_i) = p \log \lambda - p \log(\lambda + survtime_i)$ 
    llk_nt_1[i] <-
    p*log(lambda)-p*log(lambda+survtime[i])
```



## Section 6.10

```
# Likelihood when  $state_i = 1$ ,  $transplant_i = 0$ 
#  $\log(lik_i) = \log p + \log \tau + p \log \lambda$ 
#  $-(p + 1) \log(\lambda + timetotransplant_i + \tau \times survtime_i)$ 
llk_t_0[i] <-
log(p)+log(tau)+p*log(lambda)
-(p+1)*log(lambda+timetotransplant[i]+tau*survtime[i])
# Likelihood when  $state_i = 1$ ,  $transplant_i = 1$ 
#  $\log(lik_i) = p \log \lambda$ 
#  $-p \log(\lambda + timetotransplant_i + \tau \times survtime_i)$ 
llk_t_1[i] <- p*log(lambda)
-p*log(lambda+timetotransplant[i]+tau*survtime[i])
```

## Section 6.10

```
# Likelihood:
loglk[i] <-
  (1-transplant[i])*(1-state[i])*llk_nt_0[i]
  +(1-transplant[i])* state[i] *llk_nt_1[i]
  + transplant[i] *(1-state[i])*llk_t_0[i]
  + transplant[i] * state[i] *llk_t_1[i]
# 0's trick. Define  $\mu_i = -\log(lk_i) + 10000 > 0$ 
mu[i] <- -loglk[i]+10000
# Define an auxiliary variable  $zeros_i \equiv 0$ 
zeros[i] <- 0
#  $zeros_i \sim \text{Poisson}(\mu_i)$ ,  $Pr(zeros_i = 0) = lk_i * \exp(-10000)$ 
zeros[i] ~ dpois(mu[i])
# the end of for loop
}
```

## Section 6.10

```
#  $\tau = \exp(\log \tau)$ ,  $\lambda = \exp(\log \lambda)$ ,  $p = \exp(\log p)$ 
tau <- exp(ltau)
lambda <- exp(llambda)
p <- exp(lp)
# Prior:  $\log \tau \sim N(0, 10^6)$ 
ltau ~ dnorm(0, 0.000001)
# Prior:  $\log \lambda \sim N(0, 10^6)$ 
llambda ~ dnorm(0, 0.000001)
# Prior:  $\log p \sim N(0, 10^6)$ 
lp ~ dnorm(0, 0.000001)
```

## Section 6.10

```
# Estimate  $Pr(\tau \leq 1)$  using  $\text{tau.le.1} = I(\tau \leq 1)$ 
tau.le.1 <- step(1 - tau)
#  $S(t)$ : survival function.  $t = 1, \dots, 240$ 
for(t in 1:240)
{
  S[t] <- exp( p*log(lambda) - p*log(lambda + t) )
  time[t] <- t
}
# the end of the model statement
}
# set initial values for MCMC
list( ltau = 0, llambda=0, lp=0 )
# define variables for the dataset
list(n = 82)
```

## Section 6.10

```
# define variables:  survtime, transplant,  
timetotransplant, state  
survtime[] transplant[] timetotransplant[] state[]  
49 0 0 0  
...  
43 1 5 1  
END
```

## **7 Hierarchical modeling**

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## Section 7.4 Equal mortality rates

```
# First we specify the statistical model
model
{
  for(i in 1:n)
  {
    #  $y_i | \lambda \sim POI(e_i \times \lambda)$ 
    y[i] ~ dpois(e.lambda[i])
    e.lambda[i] <- e[i]*lambda
  } # Prior:  $\lambda \sim Gamma(0.001, 0.001)$ 
  lambda ~ dgamma(0.001, 0.001)
```

## Section 7.4 Equal mortality rates

```
# Posterior predictive analysis
#  $Pr(Y_{i,pred} < Y_i | data) < 0.05?$ 
# or  $Pr(Y_{i,pred} > Y_i | data) < 0.05?$ 
# If yes, the model may not be valid for  $Y_i$ .
for(i in 1:n)
{
ypred[i] ~ dpois(e.lambda[i])
pright[i] <- step(ypred[i] - y[i])
pleft[i] <- step(y[i] - ypred[i])
}
# Initial values.
list( lambda = 1 )
# Initialize other values using "gen inits".
```



## Section 7.4 Equal mortality rates

```
# Data
list(n=94)
# Data
e[] y[]
532 0
... ..
12131 17
END
```

## Section 7.5 Modeling a prior belief of exchangeability

```
# First we specify the statistical model
model
{
  for(i in 1:n)
  {
    #  $y_i | \lambda_i \sim \text{POI}(e_i \times \lambda_i)$ 
    #  $\lambda_i \sim \text{Gamma}(\alpha, \alpha/\mu)$ ,  $E(\lambda_i) = \mu$ ,  $\text{Var}(\lambda_i) = \mu^2/\alpha$ 
    y[i] ~ dpois(e.lambda[i])
    e.lambda[i] <- e[i]*lambda[i]
    lambda[i] ~ dgamma(alpha, beta)
    loge[i] <- log(e[i])
  }
  beta <- alpha*mu.inv
}
```

## Section 7.5 Modeling a prior belief of exchangeability

```
# Prior:  $\mu \sim IG(0.001, 0.001)$ 
mu.inv ~ dgamma(0.001, 0.001)
mu <- 1/mu.inv
# Prior:  $\alpha \sim g(\alpha) = z_0/(\alpha + z_0)^2$ ,  $z_0 = 0.53$ .
#  $z_0$  is the median of  $\alpha$ .
alpha <- x-z0
z0 <- 0.53
#  $f(x) = z_0/x^2, x > z_0, (x = \alpha + z_0)$ .
x ~ dpar(1, z0)
lalpha <- log(alpha)
lmu <- log(mu)
```

## Section 7.5 Modeling a prior belief of exchangeability

```
# Compute shrinkage size  $B_i$  toward the pooled
# estimate  $\mu$ :  $E(\lambda_i|y, \alpha, \mu) = (1 - B_i)\frac{y_i}{e_i} + B_i\mu$ 
# where  $B_i = \alpha/(\alpha + e_i\mu)$ 
for(i in 1:n)
{
  B[i] <- alpha/(alpha+e[i]*mu)
}
# Posterior predictive analysis
for(i in 1:n)
{
  ypred[i] ~ dpois(e.lambda[i])
  pright[i] <- step(ypred[i] - y[i])
  pleft[i] <- step(y[i] - ypred[i])
}
```

## Section 7.5 Modeling a prior belief of exchangeability

```
# Initial values.
list( mu.inv=1, x=1 )
# Initialize other values using "gen inits".      #
Data
list(n=94)
# Data
e[] y[]
532 0
... ..
12131 17
END
```

## 9.2 Regression model

---

## Section 9.2

```
# First we specify the statistical model
model
{
  for(i in 1:n)
  {
    #  $\log \text{time}_i \sim N(\mu, \sigma^2)$ 
    # where  $\mu_i = \beta_1 + \beta_2 \times \text{nesting}_i + \beta_3 \times \text{size}_i + \beta_4 \times \text{status}_i$ 
    logtime[i] <- log(time[i])
    logtime[i] ~ dnorm(mu[i], inv.sigma2)
    mu[i] <- b[1] + b[2]*nesting[i] + b[3]*size[i] +
    b[4]*status[i]
  }
}
```

## Section 9.2

```
# Prior:  $\beta_i \sim N(0, 10^6)$ ,  $i = 1, 2, 3, 4$ 
b[1] ~ dnorm(0, 0.000001)
b[2] ~ dnorm(0, 0.000001)
b[3] ~ dnorm(0, 0.000001)
b[4] ~ dnorm(0, 0.000001)
# Prior:  $\sigma^{-2} \sim IG(10^{-4}, 10^{-4})$ 
inv.sigma2 ~ dgamma(0.0001, 0.0001)
sigma2 <- 1/inv.sigma2
sigma <- sqrt(sigma2)
```



## Section 9.2

```
# Prediction of the new mean  $\mu_{pi}$  and
# the new observation  $y_{pi} \sim N(\mu_{pi}, \sigma^2)$ ,  $i = 1, 2, 3, 4$ 
for(i in 1:4)
{
mu_p[i] <-b[1]+b[2]*nesting_p[i]+b[3]*size_p[i]
+b[4]*status_p[i]
y_p[i] ~ dnorm(mu_p[i], inv.sigma2)
}
# Prediction of the new observation  $y_i^* \sim N(\mu_i, \sigma^2)$ 
for(i in 1:n)
{
ystar[i] ~ dnorm(mu[i], inv.sigma2)
# Compute  $pright_i = I(ystar_i > \log time_i)$ 
# to estimate  $Pr(ystar_i > \log time_i)$ 
```

## Section 9.2

```
pright[i] <- step(ystar[i] - logtime[i])
# Compute  $p_{left_i} = I(ystar_i < \log time_i)$ 
# to estimate  $Pr(ystar_i < \log time_i)$ 
pleft[i] <- step(logtime[i] - ystar[i])
}
# the end of the model statement
}
# define variables for the dataset
list(n = 62)
```

## Section 9.2

```
# define variables:
time[] nesting[] size[] status[]
3.030 1.000 0 1
...
1.000 1.000 1 1
END
nesting_p[] size_p[] status_p[]
4 0 0
...
4 1 1
END
# set initial values for MCMC
list( b=c(0, 0, 0, 0), inv.sigma2 = 1, y_p=c( 0,
0, 0, 0))
```

## 11.4 Change point model

---

## Section 11.4 Change point model

```
# First we specify the statistical model
model
#  $D_t \sim POI(\mu_t)$ ,  $\mu_t = \exp(\beta_1 + I(t \geq t_0) * \beta_2)$ ,
#  $t_0$ : change point
#  $\mu_t = \exp(\beta_1)$  for  $t < t_0$ ,  $\mu_t = \exp(\beta_1 + \beta_2)$  for  $t \geq t_0$ .
{
for( year in 1 : N )
{
D[year] ~ dpois(mu[year])
log(mu[year]) <- b[1] + step(year - changeyear) * b[2]
}
}
```

## Section 11.4 Change point model

```
# Prior:  $\beta_1 \sim N(0, 10^6), \beta_2 \sim N(0, 10^6)$ 
b[1] ~ dnorm( 0.0, 1.0E-6)
b[2] ~ dnorm( 0.0, 1.0E-6)
# Prior:  $Pr(t_0 = k) = 1/N, k = 1, \dots, N.$ 
for( year in 1 : N ) p[year] <- 1/N )
changeyear ~ dcat(p[])
# Initial values
list(b = c(0, 0), changeyear = 50)
# Data
list( N = 112, D=c(4,5,4,1,0,4,3,4,0,6,...,0,0))
```